Early-phase Dose-finding Design for Molecularly Targeted Agents

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Outline

- Background
- Models & inference
- Dose-finding algorithm
- Simulation results
- Conclusions

Traditional Dose-finding

- Primary goal for a cytotoxic agent is to identify MTD
- Recommended dose further investigated in Phase II to evaluate efficacy
- Common assumption for cytotoxics is both efficacy and toxicity increase monotonically with dose
- Recent development of molecularly targeted agents (MTA's) challenges this assumption

Dose-finding For Molecularly Targeted Agents

- Toxicity may be minimal
 - Often reasonable to assume monotonicity
- Dose-efficacy curves may be non-monotonic
- Goal is to find optimal biological dose (OBD)
 - Defined by dose with acceptable toxicity that maximizes efficacy

Published Methods For MTA's

- Hunsberger, Rubinstein, Dancey, Korn; Statist Med 2005
- Mandrekar, Cui, Sargent; Statist Med 2007
- Polley, Cheung; *Biometrics* 2008
- Hoering, LeBlanc, Crowley; Clin Cancer Res 2011
- Hoering, Mitchell, LeBlanc, Crowley; Clin Trials 2013
- > Yin, Zheng, Xu; *Statist Med* 2013

Non-monotone Dose-efficacy Curves

- May exhibit unimodal or plateau relationships
 - Increase initially, then decrease or remain constant
- Suppose we have a set of *I* doses, {*d*₁,...,*d_I*} and probability of efficacy at *d_i* is denoted, π_E(*d_i*)
- ▶ The goal is to find OBD, $d_{\nu} \in \{d_1, \ldots, d_I\}$, defined by

$$\pi_E(d_1) \leq \cdots \leq \pi_E(d_\nu) \geq \cdots \geq \pi_E(d_I).$$

Possible Dose-toxicity/efficacy Curves

Yin, Zheng, Xu; Statist Med 2013



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Example

- Phase I study of Bavituximab in patients with advanced solid tumors (Gerber et al.; Clinical Cancer Research, 2011)
- ► Four available dose levels (0.1, 0.3, 1, 3 mg/kg)
- Goal of the study was dose escalation to the OBD, rather than MTD
- "... for monoclonal antibodies, the MTD may not correspond to optimal efficacy..."

Overall Strategy

- d_v corresponds to the peak of unimodal curve or beginning of plateau
- Peak could occur at any of the I available levels
 - Non-decreasing probabilities before the peak, non-increasing probabilities after peak
- Overall strategy is to formulate a set of possible dose-efficacy relationships corresponding to various "peak" locations

Possible Dose-efficacy Curves

- 1. monotone decreasing: peak at d_1 $\pi_E(d_4) \le \pi_E(d_3) \le \pi_E(d_2) \le \pi_E(d_1)$
- 2. monotone increasing: peak at d_4 $\pi_E(d_1) \le \pi_E(d_2) \le \pi_E(d_3) \le \pi_E(d_4)$.
- 3. unimodal or plateau: peak at d_2 $\pi_E(d_1) \le \pi_E(d_3) \le \pi_E(d_4) \le \pi_E(d_2)$
- 4. unimodal or plateau: peak at d_3 $\pi_E(d_1) \le \pi_E(d_2) \le \pi_E(d_4) \le \pi_E(d_3).$

Modeling Toxicity - CRM

• The probability of DLT at dose level d_i is

$$\pi_T(d_i) \approx F(d_i, \beta) = p_i^{\exp(\beta)}$$

After j inclusions, DLT probability estimates

$$\widehat{\pi}_T(d_i) = F(d_i, \widehat{\beta}_j)$$

define an acceptable set of doses based on a maximum acceptable toxicity rate ϕ_T

Models For Efficacy

- Suppose there are K dose-efficacy model possibilities under investigation.
- ► For a particular model, k, the probability of efficacy modeled by

$$\pi_E(d_i) \approx G_k(d_i, \theta) = q_{ik}^{\exp(\theta)}$$

for a class of working dose-efficacy models, $G_k(d_i, \theta)$

▶ $0 < q_{1k} < \cdots < q_{Ik} < 1$ represents the skeleton of model *k*.

Inference

- Prior probabilities on each model τ(k) = {τ(1),...,τ(K)} and prior distribution, g(θ), on θ
- After inclusion of the first *j* patients into the study, the likelihood under model *k* is given by

$$L_k(\theta \mid \mathcal{D}_j) = \prod_{\ell=1}^{j} \{G_k(x_{\ell}, \theta)\}^{z_{\ell}} \{1 - G_k(x_{\ell}, \theta)\}^{(1-z_{\ell})}$$

which, for each model, can be used in order to generate the posterior mean, $\hat{\theta}_{ik}$, of parameter θ .

Model Selection

The posterior probabilities of the models given the data are

$$\omega(k \mid \mathcal{D}_j) = \frac{\tau(k) \int L_k(\theta \mid \mathcal{D}_j) g(\theta) d\theta}{\sum_{k=1}^{K} \tau(k) \int L_k(\theta \mid \mathcal{D}_j) g(\theta) d\theta}$$

When a new patient is to be enrolled, we choose a single model, k*, with the largest posterior probability such that

$$k^* = \arg\max_k \omega(k \mid \mathcal{D}_j)$$

Model Selection in CRM

- Bayesian model averaging CRM (Yin, Yuan; JASA 2009)
- Extended model-based designs for more complex dose-finding studies (O'Quigley, Conaway; *Statist Med* 2011)
- Posterior maximization and averaging for Bayesian working model choice in the CRM (Daimon, Zohar, O'Quigley; *Statist Med* 2011)
- CRM for partial ordering (Wages, Conaway, O'Quigley; *Biometrics* 2011)

Efficacy Probability Estimates

- ► Take the working model, $G_{k^*}(d_i, \theta)$, associated with k^* to generate efficacy probability estimates at each dose.
- Compute the posterior probability of a response for d_i

$$\widehat{\pi}_E(d_i) = G_{k^*}(d_i, \widehat{\theta}_{jk^*})$$

from which we can make decisions regarding allocation.

Dose-finding Algorithm

- Overall, allocate the next entered patient to the dose estimated to be the most efficacious, among those with acceptable toxicity.
- Define the set of "acceptable" doses as

$$\mathcal{A}_i = \{ d_i : \, \widehat{\pi}_T(d_i) \le \phi_T; \, i = 1, \dots, I \}.$$

The allocation algorithm depends upon the amount of data that has been observed so far in the trial.

Randomization Phase

For doses in A_i , calculate a randomization probability R_i ,

$$R_i = \frac{\widehat{\pi}_E(d_i)}{\sum_{d_i \in \mathcal{A}_i} \widehat{\pi}_E(d_i)}$$

and randomize the next patient or cohort of patients to dose d_i with probability R_i .

Switch to a phase in which we simply allocate according to the maximum estimated efficacy probability among the acceptable doses.

Maximization Phase

► Among the doses contained in A_j, we allocate the (j + 1)th patient cohort to the dose x_{j+1} according to the estimated efficacy probabilities, \$\overline{\alpha_E}(d_i)\$, such that

$$x_{j+1} = \arg\max_{d_i \in \mathcal{A}_j} \widehat{\pi}_E(d_i)$$

► The optimal dose is the recommended dose d_i = x_{n+1} for the hypothetical (n + 1)th patient after the inclusion of the maximum sample size of n patients.

Design Specifications

- Skeleton for toxicity, $p_i = \{0.01, 0.08, 0.15, 0.22, 0.29, 0.36\}$
- Define maximum acceptable toxicity rate $\phi_T = 0.33$
- Total sample size N = 48
- ▶ Randomization phase sample size $n_R = 12$
- Each of the k models has Normal prior distribution on θ

 $g(\theta) \sim \mathcal{N}(0, 1.34)$

Skeletons For Unimodal Relationships

- peak at d₆[k = 1] {0.10, 0.20, 0.30, 0.40, 0.50, 0.60}
- ▶ peak at d₅[k = 2] {0.20, 0.30, 0.40, 0.50, 0.60, 0.50}
- ▶ peak at d₄[k = 3] {0.30, 0.40, 0.50, 0.60, 0.50, 0.40}
- ▶ peak at d₃[k = 4] {0.40, 0.50, 0.60, 0.50, 0.40, 0.30}
- ▶ peak at d₂[k = 5] {0.50, 0.60, 0.50, 0.40, 0.30, 0.20}
- ▶ peak at d₁[k = 6] {0.60, 0.50, 0.40, 0.30, 0.20, 0.10}

Skeletons For Plateau Relationships

- plateau at d₁[k = 7] {0.60, 0.60, 0.60, 0.60, 0.60, 0.60}
- ▶ plateau at d₂[k = 8] {0.50, 0.60, 0.60, 0.60, 0.60, 0.60}
- ▶ plateau at $d_3[k = 9]$ {0.40, 0.50, 0.60, 0.60, 0.60, 0.60}
- ▶ plateau at $d_4[k = 10]$ {0.30, 0.40, 0.50, 0.60, 0.60, 0.60}
- plateau at d₅[k = 11] {0.20, 0.30, 0.40, 0.50, 0.60, 0.60}
- Assume, a priori, that each skeleton is equally likely and set $\tau(k) = 1/11$

Association Between Toxicity & Efficacy

- Association between toxicity and efficacy is ignored in modeling
- Simulation studies assess the sensitivity of method to association between toxicity and efficacy
- \blacktriangleright Correlated binary responses generated under various values of association parameter ψ
- Results presented with $\psi = 4.6$ as in Hoering et al. (*Clinical Trials*, 2013)

Dose-finding in Hoering et al. (2013)

- 3+3 design in Phase I to find recommend dose (RD) for Phase II
- In Phase II, randomize 48 patients to one of three arms (dose levels)
 - ▶ RD-1, RD, RD+1
- Choose dose with highest efficacy that is also safe (DLT rate ≤ 33%.
- Authors define best dose as the level that maximizes efficacy while assuring safety; good dose as level where efficacy is above predefined boundary while maintaining safety.

Hoering et al. Scenarios



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Simulation Results



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Unimodal Dose-efficacy Curves

Scenario 1 (π_T, π_E)

- $\blacktriangleright \ \{(0.05, 0.20), (0.10, 0.40), (0.25, 0.60), (0.45, 0.80), (0.60, 0.55)\}$
- Peak of d-e curve occurs outside acceptable set
- Scenario 2 (π_T, π_E)
 - $\blacktriangleright \ \{(0.08, 0.20), (0.12, 0.40), (0.20, 0.60), (0.30, 0.80), (0.42, 0.55)\}$
 - Peak of d-e curve occurs at edge of acceptable set
- Scenario 3 (π_T, π_E)
 - $\blacktriangleright \ \{(0.06, 0.20), (0.08, 0.40), (0.14, 0.60), (0.20, 0.80), (0.30, 0.55)\}$
 - Peak of d-e curve occurs inside acceptable set

Results

True DLT prob.	0.05	0.10	0.25	0.45	0.60
True Eff prob.	0.20	0.40	0.60	0.80	0.55
% selection	0.04	0.19	0.53	0.24	0.00
% allocation	0.12	0.26	0.35	0.23	0.03
True DLT prob.	0.08	0.12	0.20	0.30	0.42
True Eff prob.	0.20	0.40	0.60	0.80	0.55
% selection	0.04	0.17	0.29	0.50	0.01
% allocation	0.13	0.23	0.25	0.34	0.05
True DLT prob.	0.06	0.08	0.14	0.20	0.30
True Eff prob.	0.20	0.40	0.60	0.80	0.55
% selection	0.04	0.13	0.19	0.62	0.03
% allocation	0.10	0.19	0.20	0.43	0.07
$N = 30; n_R = 10; \phi_T = 0.35, \psi = -2$					

Conclusions

- Bivariate extension of CRM for effectively estimating optimal dose in early-phase trials of targeted agents.
- Good operating characteristics when compared to published method in area.
- Extension includes relaxing monotonicity assumption for toxicity
 - For trials of dual-agent combinations
 - Use partial order CRM (Wages, Conaway, O'Quigley; *Biometrics* 2011) to estimate DLT probabilities
- Exploring stopping rules for toxicity and efficacy
- Modifications for delayed response outcomes

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